



Early smoking-onset age and risk of cardiovascular disease and mortality

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ABSTRACT

Early smoking onset age (SOA) is a public health concern with scant empirical evidence of its role in health outcomes. The study had two aims: i) to assess whether an early SOA was associated with the risk of fatal and non-fatal CVD and all-cause and CVD mortality and ii) to explore the linear and non-linear association between SOA and the outcomes of interest. Data from 4499 current or former smokers, recruited from 1995 to 2005, aged 25 to 79 years, and with a median 7.02 years of follow-up, were obtained from the REGICOR population-based cohort. In the present analysis, performed in 2018, the independent variable was SOA and the dependent variables were CVD events, CVD mortality, and all-cause mortality. Penalized smoothing spline methods were used to assess the linear and non-linear association. During follow-up, 361 deaths and 210 CVD events were recorded. A significant non-linear component was identified in the association between SOA and CVD outcomes with a cut-off point at 12 years: In the group aged ≤ 12 years, each year of delay in SOA was inversely associated with CVD risk (HR = 0.71; 95%CI = 0.53–0.96) and CVD mortality (HR = 0.58; 95%CI = 0.37–0.90). No association was observed in the older SOA group. A linear association was observed between SOA and all-cause mortality, and each year of delay was associated with 4% lower risk of mortality (HR = 0.96; 95%CI = 0.93–0.98). The associations were adjusted for lifelong exposure to tobacco and cardiovascular risk factors. These results reinforce the value of preventing tobacco use among teenagers and adolescents.

1. Introduction

Consistent evidence supports the role of smoking as a risk factor for cardiovascular disease (CVD) (Gordon and Flanagan, 2016; Ockene and Miller, 1997; McEvoy et al., 2015). Smokers have about twice the risk of coronary heart disease (Stallones, 2015) and stroke (Kelly et al., 2008), compared to non-smokers. Smoking is also a risk factor for ischemic nephropathy (Criqui and Aboyans, 2015), bowel ischemia (Tang et al., 2016) aortic dissection (Shiraishi et al., 2014), cancer and all-cause mortality (Hohenwarter, 2009). Furthermore, smoking has been described as nearly a prerequisite for the development of peripheral arterial disease (Landenhed et al., 2015) and abdominal aortic

aneurysms (Doll et al., 2004).

Several smoking indicators have been explored and analysed, with particular attention to smoking status (current, former and never smokers) and cumulative exposure to tobacco (Wu et al., 2013; Mons et al., 2015). Early smoking-onset age (SOA) is another important indicator of exposure, as 68.1% of smokers in Europe start before 18 years of age and the mean age for onset of regular smoking is 16.6 years (Dani and Harris, 2005). An early SOA has been associated with psychiatric disorders (Genuneit et al., 2006), asthma (Wiencke and Kelsey, 2002), lung cancer and other malignancies (Peto et al., 2000; Funatogawa et al., 2012; Choi and Stommel, 2017; Catsburg et al., 2015; Park et al., 2014; Kenfield et al., 2008; Huxley et al., 2012) and all-cause mortality

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(Wu et al., 2013; Catsburg et al., 2015; Huxley et al., 2012). Although some studies suggest a positive association between early SOA and CVD (Catsburg et al., 2015; Planas et al., 2002; Honjo et al., 2010; Filippidis et al., 2015), potential limitations include a lack of adjustment for cumulative tobacco exposure (Filippidis et al., 2015) or vascular risk factors (Catsburg et al., 2015; Planas et al., 2002; Filippidis et al., 2015), a retrospective design (Catsburg et al., 2015; Honjo et al., 2010) and under-representation by sex (Wu et al., 2013; Huxley et al., 2012; Honjo et al., 2010). Moreover, the pattern of dose-response relationships between SOA and health outcomes has not been fully explored.

The study had two aims: i) to assess whether an early SOA was associated with the risk of fatal and non-fatal CVD and all-cause and CVD mortality and ii) to explore the linear and non-linear association between SOA and the clinical outcomes of interest in a Mediterranean population in southern Europe.

2. Material and methods

2.1. Study design and population

The REGICOR (REgistre Gironi del COR, or Girona Heart Registry) study recruited a prospective population-based cohort in Girona province (~700,000 inhabitants) in northeastern Spain, with the objective of studying CVD and related risk factors over time. Recruitment details have been described elsewhere (Grau et al., 2007). Briefly, individuals living in 42 communities, including 41 villages and the city of Girona, were randomly selected from the census and invited to participate. Inclusion criteria required that participants were free of terminal disease, not institutionalized, and had lived in the referral area for at least six months/year (reflecting the stable seasonal presence of a large number of retirees).

Participants were recruited for three different surveys: 1748 residents aged 25 to 74 years in 1995, 3058 aged 25 to 74 years in 2000, and 6352 aged 35 to 79 years in 2005. Selected participants received a letter informing them of the overall aims of the study, the purpose of the specific survey, and the tests to be performed. The participation rates in these surveys were 72.4%, 70.0% and 73.8%, respectively.

The present analysis included participants from all three surveys. In the case of participants in more than one survey, the longest available follow-up data were considered. We excluded those participants with personal history of CVD and those older than 79 years at the date of study inclusion, or reporting never having smoked. As very few smokers started before 9 years or after 30 years of age, these individuals were also excluded from analysis. From an initial sample of 11,158 individuals, 4499 were finally included in the main analysis as shown in flow-chart (Fig. 1). The study protocol was approved by the Parc de Salut Mar Research Ethics Committee (2008/3046/I; 2016/7075/I) and each participant signed an informed consent.

2.2. Baseline data

Participants were asked to fast for at least 10 h before their appointment at the examination site, which included drawing a blood sample. A group of nurses trained in the study protocol administered a set of validated and standardized questionnaires (Manual of the MONICA Project, n.d.; Baena-Díez et al., 2009) and performed a physical examination.

Self-reported educational level (elementary school, secondary school or university degree) was considered as an indicator of socioeconomic position. Hypertension was defined if the individual was previously diagnosed by a physician or receiving treatment or presented with systolic blood pressure (SBP) ≥ 140 mmHg or diastolic blood pressure (DBP) values ≥ 90 mmHg. Total cholesterol, high density lipoprotein (HDL) cholesterol, triglycerides and glucose concentrations were determined by enzymatic methods. When triglycerides were < 300 mg/dL, low density lipoprotein (LDL) cholesterol was

estimated by the Friedewald formula. Diabetes was defined if previously diagnosed or treated or when an individual presented with fasting glucose values ≥ 126 mg/dL.

2.3. Smoking data

Smoking exposure was assessed by a standardized questionnaire (Manual of the MONICA Project, n.d.; Baena-Díez et al., 2009). Participants were classified as smokers (current or quit < 1 year) or former smokers (quit > 1 year). SOA was defined as the reported age at which the individual regularly smoked at least 1 cigarette/day. Lifetime pack-years was estimated as years of smoking multiplied by number of 20-cigarette packs consumed daily. Cigar smoking was included in the analysis according to tobacco equivalence to cigarettes (Kozłowski et al., 2008). Pipe smoking was negligible within our population and was not considered.

Among current smokers (51.9% of participants), years of smoking were quantified as the difference between age at the REGICOR survey visit and SOA. Among former smokers (48.1% of participants), smoking exposure time was calculated as the difference between cessation age and SOA. As the type of information available about smoking cessation date varied between the three REGICOR surveys, the exact year former smokers quit smoking was available for 35.2% of participants. In those with missing data, we used multiple imputation methods implemented in the mice R package and obtained 20 data sets to impute the value of this variable. To obtain an estimator of the associations of interest we used the MIcombine function in R (R Core Team, 2018).

2.4. Follow-up and outcomes

The follow-up included physical re-exams and contact by telephone every two years. The re-exams included physical examination and all previously administered questionnaires. In those not attending the physical re-exams, a follow-up telephone contact was attempted to identify the appearance of events of interest, by means of a standard questionnaire. To ascertain any cardiovascular events or deaths, we also reviewed medical records, linked the data with a population-based myocardial infarction register, and cross-checked all these sources of information. To identify fatal events not otherwise reported, we linked our data with the regional mortality register (until 31 December 2010).

Three main outcomes were defined: i) CVD events, including non-fatal myocardial infarction or stroke and fatal CVD events (ICD9 codes: 390-459, 798 or ICD10 codes: I00-I99, R96, R98-99), ii) CVD mortality, and iii) all-cause mortality. All the events were classified by an event committee according to standardized criteria (Velescu et al., 2017). In case of multiple CVD events in the same participant, the first occurring event was considered in defining the composite CVD outcome.

2.5. Statistical analysis

The analysis was performed in 2018. Continuous variables were expressed as mean and standard deviation or median and interquartile range, and categorical data as frequencies and percentages. Categorical variables were compared with the Chi-square or Fisher exact test, as appropriate, and continuous variables with the *t*-test or ANOVA for normal distribution and the Mann-Whitney *U* test or Kruskal-Wallis test for non-normal distributed variables.

To explore the hypothesis that an early SOA was independently associated with an increased risk of CVD event or mortality, we used penalized smoothing spline (pspline function, R Survival Package) which allows a maximum of 10 knot points, to assess the linear and non-linear components of the dose-response association (Eilers and Marx, 1996). When the non-linear component was nonsignificant, SOA was considered as a continuous variable in Cox proportional hazard regression models. Otherwise, we used bootstrapping methods to define the best cut-point(s) at which a change in the linear dose-response

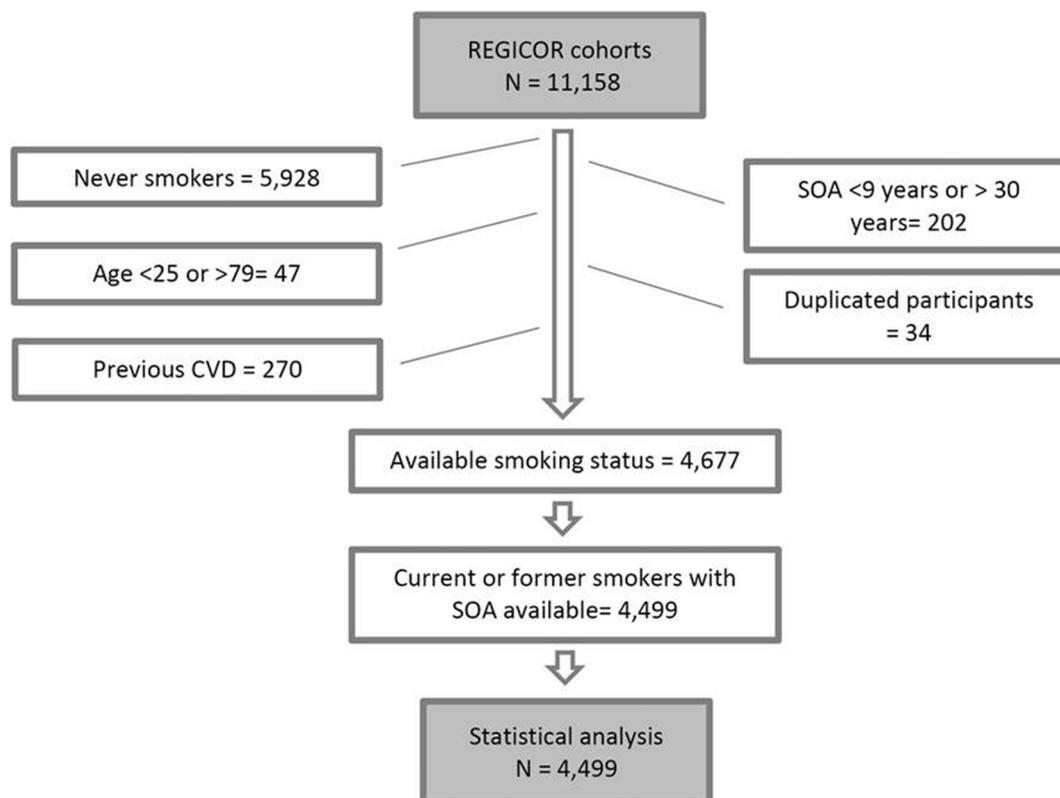


Fig. 1. Flow chart of participant selection and inclusion. REGICOR population-based cohort study, participants recruited from 1995 to 2005, aged 25 to 79 years, with a median 7.02 years of follow-up.

association was observed. We performed 1000 iterations per outcome of interest and calculated the median of the observed cut-point, which was then defined as the best cut-point. The analysis was stratified according to best cut-point, and Cox regression modelling considered SOA as a continuous variable in each of the defined strata.

The assumption of the proportionality of risks according to the Schoenfeld residuals was checked in each Cox regression model. To control for potential confounding factors, all variables related to the outcomes of interest with a $p < 0.05$ in bivariate analyses were considered in the multivariate model. Statistically non-significant variables were removed step-by-step from the model using a backward procedure when non-significant variations occurred in the regression coefficients of the SOA. Lifetime smoking exposure (pack-years), smoking status (current vs former smokers), years since quitting smoking, and recruitment survey were included in all models regardless of significance. The statistical analysis adopted a competing risk strategy using the Gray method, considering non-CVD death causes as the competing event for a CVD event or CVD-related mortality, and other death causes as the competing event for CVD-related mortality. As a sensitivity analysis, the results were stratified by current smoking status and by sex.

A p -value < 0.05 was considered as statistically significant. All analyses were performed using the R statistical package (R Core Team, 2018).

3. Results

Participant characteristics are summarized according to SOA groups (stratified as ≤ 12 y and > 12 y, based on the findings reported below). Earlier SOA was more likely in men and was associated with a higher number of pack-years. In the bivariate analysis, an earlier SOA was also associated with increased risk of CVD events, CVD mortality, and all-cause mortality (Table 1).

3.1. Cardiovascular events and smoking-onset age

During follow-up (median 7.02 years), 210 fatal and non-fatal CVD events were recorded. The non-linear component of the association between SOA and CVD events was statistically significant (p -value = 0.002) (Fig. 2, panel A). The bootstrapping analyses showed that the best cut-point to define a change in the linear dose-response association was 11–12 years for cardiovascular events (12 years for cardiovascular mortality). Therefore, the analysis was stratified in two SOA groups: ≤ 12 y and > 12 y. In the younger group (SOA from 9 to 12 years), each year of delay in SOA was associated with a 29% decrease in CVD event risk (HR = 0.71; 95% Confidence Interval-CI: 0.53–0.96) (Table 2). In the older group (SOA from 13 to 30 years), SOA was not associated with CVD risk (HR = 1.00; 95% CI: 0.96–1.04). Similar results were observed in the sensitivity analysis in current and former smokers (Appendix Table A1–A2 and Fig. A1–A2) and in men (Appendix Table A3). In women, the models in the younger SOA age group did not converge due to the low number of events (Appendix Table A4).

3.2. Mortality and smoking onset age

During follow-up, there were 361 deaths (77 CVD-related). The non-linear component of the association between SOA and CVD mortality was statistically significant (p -value = 0.010) and was non-significant for all-cause mortality (p -value = 0.220) (Fig. 2, panel B and C).

The bootstrapping analyses showed that the best cut-point to define a change in the linear dose-response association for cardiovascular mortality was 12 years. Therefore, the analysis was stratified in two SOA groups: ≤ 12 years and > 12 years. In the younger group, SOA was inversely associated with the risk of CVD mortality. Each year of delay in SOA was associated with a 42% decrease in CVD mortality risk (HR = 0.58; 95% CI: 0.37–0.90) (Table 2). In the older group, SOA was

Table 1

Baseline characteristics of the study population according to smoking onset age. REGICOR population-based cohort study, participants recruited from 1995 to 2005, aged 25 to 79 years, with a median 7.02 years of follow-up.

Variables	(n)	Smoking onset age ≤ 12y (n = 244)	Smoking onset age > 12y (n = 4255)
Age, years ^a	4499	57.7 (12.3)	49.9 (12.7)
Sex: women, n (%)	4499	19 (7.8%)	1389 (32.6%)
Education, n (%)	4447	–	–
University		13 (5.5%)	895 (21.3%)
Secondary school		43 (18.2%)	1263 (30.0%)
Primary school		180 (73.3%)	2053 (48.8%)
Body mass index ^a , kg/m ²	4466	27.3 (4.7)	26.8 (4.3)
Diabetes, n (%)	4427	47 (19.5%)	429 (10.2%)
Glucose serum ^a , mg/dL	4390	108 (28.9)	101 (26.0)
Hypertension, n (%)	4488	135 (56.5%)	1564 (37.5%)
SBP ^a , mmHg	4485	136 (19.5)	126 (19.7)
DBP ^a , mmHg	4474	80.4 (10.7)	78.3 (10.7)
Total cholesterol ^a , mg/dL	4386	213 (43.5)	213 (42.0)
HDL cholesterol ^a , mg/dL	4373	49.4 (13.1)	49.4 (13.1)
LDL cholesterol ^a , mg/dL	4214	139 (39.4)	140 (37.6)
Triglycerides ^b , mg/dL	4379	103 [77;143]	97 [72;136]
Smoking exposure			
Smoking status n (%)	4499	–	–
Current smoker		111 (45.5%)	2224 (52.3%)
Former smoker > 1y		133 (54.5%)	2031 (47.7%)
Pack-years ^b	3100	32.7 [9.2;55.6]	14.0 [3.6;28.5]
Years since quitting ^a	3119	5.4 (8.9)	2.9 (6.7)
Outcomes			
Fatal/non-fatal CVD event, n (%)	4499	34 (13.9%)	176 (4.1%)
CVD mortality, n (%)	4499	17 (7.0%)	60 (1.4%)
All-cause mortality, n (%)	4499	59 (24.2%)	302 (7.1%)

SBP: systolic blood pressure; DBP: diastolic blood pressure; HDL: high density lipoprotein; LDL: low density lipoprotein; CVD: cardiovascular disease.

^a Mean (standard deviation).

^b Median [interquartile range].

not associated with CVD risk (HR = 1.00; 95% CI: 0.93–1.08). The association between SOA and all-cause mortality was linear. Each year of delay in SOA was associated with 4% decrease in all-cause mortality

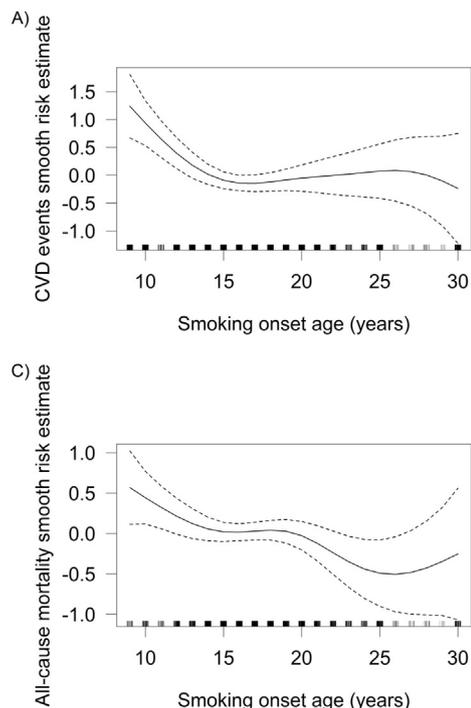


Table 2

Multivariate adjusted association between smoking onset age, as a continuous variable, and the outcomes of interest in current and former smokers. REGICOR population-based cohort study, participants recruited from 1995 to 2005, aged 25 to 79 years, with a median 7.02 years of follow-up.

	Smoking onset age	
	Smoking onset age between 9 and 12 years (n = 219)	Smoking onset age between 13 and 30 years (n = 3952)
CVD events: fatal and non-fatal		
HR	0.71 ^a	1.00 ^b
95% CI	0.53–0.96	0.96–1.04
CVD mortality		
HR	0.58 ^c	1.00 ^d
95% CI	0.37–0.90	0.93–1.08
	Smoking onset age between 9 and 30 years (n = 4171)	
All-cause mortality		
HR		0.96 ^e
95% CI		0.93–0.98

CVD: cardiovascular disease; HR: hazard ratio; CI: confidence interval.

^a Adjusted for age, total cholesterol, high-density lipoprotein cholesterol, smoking status, pack-years, recruitment survey and years since quitting smoking.

^b Adjusted for age, glucose, high-density lipoprotein cholesterol, smoking status, pack-years, recruitment survey and years since quitting smoking.

^c Adjusted for age, hypertension treatment, smoking status, pack-years, recruitment survey and years since quitting smoking.

^d Adjusted for age, glucose, smoking status, pack-years, recruitment survey and years since quitting smoking.

^e Adjusted for age, sex, glucose, smoking status, pack-years, recruitment survey and years since quitting smoking.

risk (HR = 0.96; 95% CI: 0.93–0.98) (Table 2). Similar results for both outcomes were observed in the sensitivity analysis in current and former smokers (Appendix Table A1–A2 and Fig. A1–A2) and in men (Appendix Table A3).

Fig. 2. Penalized smoothing spline plots of the linear and non-linear dose-response association between smoking onset age and the outcomes of interest: A.-Cardiovascular events; B.-Cardiovascular mortality; C.-All-cause mortality. REGICOR population-based cohort study, participants recruited from 1995 to 2005, aged 25 to 79 years, with a median 7.02 years of follow-up. The lines along the horizontal axes represent the number of individuals across smoking onset age.

4. Discussion

This study analysed the linear and non-linear dose-response relationship between SOA and three clinical outcomes: CVD events, CVD mortality and all-cause mortality. The association between SOA and CVD fatal and non-fatal events did not follow a linear dose-response association. Two clear but different patterns were observed with a cut-point at 12 years of age. Individuals who started to smoke at or before this cutpoint showed an inverse and linear association between SOA and CVD health outcomes. In this group, each year of delay in SOA was associated with a decrease of 29% and 42% in CVD events and CVD mortality risk, respectively. In contrast, we did not find a higher SOA-related CVD risk among those who started to smoke after 12 years of age. The association between SOA and all-cause mortality followed a linear pattern: each year of delay in SOA was associated with a 4% decrease in all-cause mortality. These associations were independent of lifetime cumulative exposure to tobacco.

Several studies have analysed the association between SOA and cardiovascular events and all-cause mortality (Choi and Stommel, 2017; Catsburg et al., 2015; Park et al., 2014; Kenfield et al., 2008; Huxley et al., 2012; Planas et al., 2002; Honjo et al., 2010). However, in those studies age was categorized prior to the analysis and the linear and non-linear dose-response relationship was not specifically assessed. The use of penalized smoothing splines methods allowed us to explore those patterns of association.

Our results support a clear association between SOA and CVD health outcomes. The magnitude of the association between early SOA and CVD disease risk in our study is consistent with that reported in two large American cohorts (Catsburg et al., 2015; Planas et al., 2002). Choi et al. (Catsburg et al., 2015) used data from the United States National Health Interview Survey, in which smoking and CVD clinical events are self-reported. They reported a linear association between SOA ≤ 16 years and increased risk of CVD-related events: the earlier the SOA, the higher the CVD risk. No association between SOA and CVD risk was observed in the group older than 16 years when they started smoking. In the ARIC Study (Planas et al., 2002), SOA ≤ 18 years was related to an increased risk of CVD among current smokers: again, the earlier the SOA, the higher the CVD risk and there was no association in those who started to smoke when older than 18 years. In a European population, Planas et al. also reported a positive association between peripheral arterial disease and SOA < 16 years, with the study limitations inherent to a retrospective analysis and a small sample size (Honjo et al., 2010). In Asian populations, no clear association between early SOA and CVD mortality has been reported; however, the published studies did not evaluate the impact on non-fatal CVD events (Wu et al., 2013; Filippidis et al., 2015). Finally, the Nurses' Health Study did not find increased CVD mortality among women across several SOA groups (Huxley et al., 2012). The differences between these studies could be related to the population of reference: Caucasian (Catsburg et al., 2015; Planas et al., 2002; Honjo et al., 2010) vs Asian (Wu et al., 2013; Filippidis et al., 2015), or only women (Huxley et al., 2012) vs men and women (Catsburg et al., 2015; Planas et al., 2002; Honjo et al., 2010). Moreover, the different SOA cut-points – 12 years in our study, 16 years in Choi (Catsburg et al., 2015) and Planas et al. (Honjo et al., 2010), and 18 years in the ARIC study (Planas et al., 2002) could be explained by the reliability of self-reported questionnaire data and also by the approach used to define age groups for analysis.

Similarly, the association between early SOA and increased mortality observed in our study is consistent with the findings in previous studies (Wu et al., 2013; Catsburg et al., 2015; Huxley et al., 2012). A reduced mortality risk was associated with older SOA in the follow-up of a cohort in Bangladesh (Wu et al., 2013) and of participants in the U.S. National Health Interview Surveys (Catsburg et al., 2015) and the Nurses' Health Study (Huxley et al., 2012).

Although smoking prevention is important at any age to avoid the additional risks associated with smoking exposure, the results of our

study highlight 9 to 12 years as a critical age range. Three main mechanisms have been proposed to explain the relationship between early SOA and health outcomes. First, early initiation might be assumed to lead to higher exposure due to a longer period of exposure (Gall et al., 2014). In our study, however, the results suggest that an early SOA is an independent risk factor regardless of smoking status at the time of study recruitment or cumulative lifetime exposure to tobacco smoke. Second, as childhood and adolescence are critical periods for organ development, exposure to smoking could affect tissue maturation and ability to adapt to stress, implying an increased risk of future endothelium frailty and propensity to CVD (Gall et al., 2014). Moreover, early SOA has been associated with higher risk of substance dependence in adulthood (Dani and Harris, 2005).

Smoking remains a key public health issue in Europe (Wilkins et al., 2017) and throughout the world (GBD 2016 Risk Factors Collaborators, 2017). Among the factors defining individual exposure to smoking, early SOA has recently emerged as an important variable to consider. In Spain, the average SOA is 14.6 years, affecting 8.9% of the population aged 14–18 years (Observatorio Español de la Droga y las Toxicomanías, 2016), compared to the European mean SOA of 16.6 years (Dani and Harris, 2005). This illustrates the high prevalence of an early SOA in our setting and the need to communicate its consequences and implications in ways that reach young adolescents. Although smoking regulation policies have been shown to be effective in preventing CVD (Agüero et al., 2013), the close relationship between an early SOA and health outcomes recommends the implementation of new strategies focussing on childhood and adolescence.

Our study had several limitations. First, all smoking data were obtained by self-reported questionnaires with no objective assessment of smoking exposure. Nonetheless, a recent study suggests a good correlation between self-reported data and biochemical verification (Caraballo et al., 2001). Second, a recall bias may have influenced the quality of information on past smoking exposure (SOA, daily exposure, or quitting date) among former smokers. Third, daily exposure was considered to be constant throughout the lifelong smoking period, which is unlikely to be true for all smokers. Fourth, quitting date was not collected in a consistent format across the three REGICOR surveys. The methods implemented to establish this variable in former smokers may have affected the reported impact of early SOA on the study outcomes. Finally, our study may have been underpowered, especially in women, because of the low number of CVD events and deaths compared to studies in other geographic areas. Furthermore, the population of the Mediterranean region is characterized by a low incidence of CVD events (Liyanaage et al., 2016).

Our results contribute to the evidence that the age when an individual begins to smoke is an independent risk factor for fatal and non-fatal CVD and for all-cause mortality, independent of cumulative lifelong exposure to smoking. Our results clearly reinforce the need to implement health promotion strategies against tobacco use among teenagers and adolescents, especially before they reach 13 years of age.

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.ypmed.2019.04.022>.

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Conflict of interest

The authors declare they do not have any conflict of interest to disclose.

References

- Agüero, F., Dégano, I.R., Subirana, I., et al., 2013. Impact of a partial smoke-free legislation on myocardial infarction incidence, mortality and case-fatality in a population-based registry: the REGICOR study. *PLoS One* 8, e53722.
- Baena-Díez, J.M., Alzamora-Sas, M.T., Grau, M., et al., 2009. Validity of the MONICA cardiovascular questionnaire compared with clinical records. *Gac. Sanit.* 23, 519–525.
- Caraballo, R.S., Giovino, G.A., Pechacek, T.F., Mowery, P.D., 2001. Factors associated with discrepancies between self-reports on cigarette smoking and measured serum cotinine levels among persons aged 17 years or older: third National Health and Nutrition Examination Survey, 1988–1994. *Am. J. Epidemiol.* 153, 807–814.
- Catsburg, C., Miller, A.B., Rohan, T.E., 2015. Active cigarette smoking and risk of breast cancer. *Int. J. Cancer* 136, 2204–2209.
- Choi, S.H., Stommel, M., 2017. Impact of age at smoking initiation on smoking-related morbidity and all-cause mortality. *Am. J. Prev. Med.* 53, 33–41.
- Criqui, M.H., Aboyans, V., 2015. Epidemiology of peripheral artery disease. *Circ. Res.* 116, 1509–1526.
- Dani, J.A., Harris, R.A., 2005. Nicotine addiction and comorbidity with alcohol abuse and mental illness. *Nat. Neurosci.* 8, 1465–1470.
- Doll, R., Peto, R., Boreham, J., Sutherland, I., 2004. Mortality in relation to smoking: 50 years' observations on male British doctors. *BMJ* (328:1519–0).
- Eilers, P.H., Marx, B.D., 1996. Flexible smoothing with B-splines and penalties. *Stat. Sci.* 11, 89–121.
- Filippidis, F.T., Agaku, I.T., Vardavas, C.I., 2015. The association between peer, parental influence and tobacco product features and earlier age of onset of regular smoking among adults in 27 European countries. *Eur. J. Pub. Health* 25, 814–818.
- Funatogawa, I., Funatogawa, T., Yano, E., 2012. Impacts of early smoking initiation: long-term trends of lung cancer mortality and smoking initiation from repeated cross-sectional surveys in Great Britain. *BMJ Open* 2, e001676.
- Gall, S., Huynh, Q.L., Magnussen, C.G., et al., 2014. Exposure to parental smoking in childhood or adolescence is associated with increased carotid intima-media thickness in young adults: evidence from the Cardiovascular Risk in Young Finns study and the Childhood Determinants of Adult Health Study. *Eur. Heart J.* 35, 2484–2491.
- GBD 2016 Risk Factors Collaborators, 2017. Global Health Metrics Global, regional, and national comparative risk assessment of 84 behavioural, environmental and occupational, and metabolic risks or clusters of risks, 1990–2016: a systematic analysis for the Global Burden of Disease Study 2016. *Lancet* 390, 1345–1422.
- Genuneit, J., Weinmayr, G., Radon, K., et al., 2006. Smoking and the incidence of asthma during adolescence: results of a large cohort study in Germany. *Thorax* 61, 572–578.
- Gordon, P., Flanagan, P., 2016. Smoking: a risk factor for vascular disease. *J. Vasc. Nurs.* 34, 79–86.
- Grau, M., Subirana, I., Elosua, R., et al., 2007. Trends in cardiovascular risk factor prevalence (1995–2000–2005) in northeastern Spain. *Eur. J. Cardiovasc. Prev. Rehabil.* 14, 653–659.
- Hohenwarter, E.J., 2009. Chronic mesenteric ischemia: diagnosis and treatment. *Semin. Interv. Radiol.* 26, 345–351.
- Honjo, K., Iso, H., Tsugane, S., et al., 2010. The effects of smoking and smoking cessation on mortality from cardiovascular disease among Japanese: pooled analysis of three large-scale cohort studies in Japan. *Tob. Control.* 19, 50–57.
- Huxley, R.R., Yatsuya, H., Lutsey, P.L., et al., 2012. Impact of age at smoking initiation, dosage, and time since quitting on cardiovascular disease in African Americans and Whites: the atherosclerosis risk in communities study. *Am. J. Epidemiol.* 175, 816–826.
- Kelly, T.N., Gu, D., Chen, J., et al., 2008. Cigarette smoking and risk of stroke in the Chinese adult population. *Stroke* 39, 1688–1693.
- Kenfield, S.A., Stampfer, M.J., Rosner, B.A., Colditz, G.A., 2008. Smoking and smoking cessation in relation to mortality in women. *JAMA* 299, 2037.
- Kozlowski, L.T., Dollar, K.M., Giovino, G.A., 2008. Cigar/cigarillo surveillance. *Am. J. Prev. Med.* 34, 424–426.
- Landenhed, M., Engström, G., Gottsäter, A., et al., 2015. Risk profiles for aortic dissection and ruptured or surgically treated aneurysms: a prospective cohort study. *J. Am. Heart Assoc.* 4, e001513.
- Liyanage, T., Ninomiya, T., Wang, A., et al., 2016. Effects of the Mediterranean diet on cardiovascular outcomes—a systematic review and meta-analysis. *PLoS One* 11, e0159252.
- Manual of the MONICA Project. Geneva: World Health Organization. Available at: <http://www.thl.fi/publications/monica/index.html>. Accessed June 20, 2017.
- McEvoy, J.W., Nasir, K., DeFilippis, A.P., et al., 2015. Relationship of cigarette smoking with inflammation and subclinical vascular disease significance. *Arterioscler. Thromb. Vasc. Biol.* 35, 1002–1010.
- Mons, U., Müezzlin, A., Gellert, C., et al., 2015. Impact of smoking and smoking cessation on cardiovascular events and mortality among older adults: meta-analysis of individual participant data from prospective cohort studies of the CHANCES consortium. *BMJ* 350, h1551.
- Observatorio Español de la Droga y las Toxicomanías, 2016. Informe. In: Alcohol, tabaco y drogas ilegales en España. Plan Nacional sobre Drogas, Madrid.
- Ockene, I.S., Miller, N.H., 1997. Cigarette smoking, cardiovascular disease, and stroke. *Circulation* 96, 3243–3247.
- Park, S., Jee, S.H., Shin, H.-R., et al., 2014. Attributable fraction of tobacco smoking on cancer using population-based nationwide cancer incidence and mortality data in Korea. *BMC Cancer* 14, 406.
- Peto, R., Darby, S., Deo, H., et al., 2000. Smoking, smoking cessation, and lung cancer in the UK since 1950: combination of national statistics with two case-control studies. *BMJ* 321, 323–329.
- Planas, A., Clará, A., Marrugat, J., et al., 2002. Age at onset of smoking is an independent risk factor in peripheral artery disease development. *J. Vasc. Surg.* 35, 506–509.
- R Core Team, 2018. R: a language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria URL: <https://www.R-project.org/>.
- Shiraishi, N., Kitamura, K., Kohda, Y., et al., 2014. Prevalence and risk factor analysis of nephrosclerosis and ischemic nephropathy in the Japanese general population. *Clin. Exp. Nephrol.* 18, 461–468.
- Stallones, R.A., 2015. The association between tobacco smoking and coronary heart disease. *Int. J. Epidemiol.* 44, 735–743.
- Tang, W., Yao, L., Roetker, N.S., et al., 2016. Lifetime risk and risk factors for abdominal aortic aneurysm in a 24-year prospective study: the ARIC study (atherosclerosis risk in communities). *Arterioscler. Thromb. Vasc. Biol.* 36, 2468–2477.
- Velescu, A., Clara, A., Martí, R., et al., 2017. Abnormally high ankle-brachial index is associated with all-cause and cardiovascular mortality: the REGICOR study. *Eur. J. Vasc. Endovasc. Surg.* 54, 370–377.
- Wiencke, J.K., Kelsey, K.T., 2002. Teen smoking, field cancerization, and a “critical period” hypothesis for lung cancer susceptibility. *Environ. Health Perspect.* 110, 555–558.
- Wilkins, E., Wilson, L., Wickramasinghe, K., et al., 2017. European Cardiovascular Disease Statistics. European Heart Network, Brussels.
- Wu, F., Chen, Y., Parvez, F., et al., 2013. A prospective study of tobacco smoking and mortality in Bangladesh. *PLoS One* 8, e58516.